Review

# Efficacy of orlistat for the treatment of metabolic dysfunction-associated steatotic liver disease patients: A systematic review and meta-analysis

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# Abstract

Metabolic dysfunction-associated steatotic liver disease (MASLD) is a persistent hepatic condition linked with cardiovascular disorders and metabolic disturbances. Characterized by inflammation, fat accumulation, and fibrosis within the liver, MASLD can develop into liver cancer and cirrhosis. With a global prevalence of 32.4%, the condition parallels rising obesity rates. Orlistat inhibits lipase enzymes and, therefore, reduces dietary fat absorption, which may benefit MASLD patients. The present systematic review and meta-analysis were performed in accordance with PRISMA guidelines. Searches of PubMed, Scopus, Web of Science, and Embase up to January 2025 were performed using specific keywords and MeSH terms. Bias assessment and data extraction were conducted using Joanna Briggs Institute (JBI) tools independently by two researchers. Statistical analyses were performed with Stata version 14, calculating standardized mean differences, 95% confidence intervals (CI), and heterogeneity (by performing Cochran's Q test and I<sup>2</sup> index). Moreover, meta-regression, subgroup analyses, and sensitivity analyses were conducted. Eleven studies featuring 582 participants were included. Orlistat treatment induced a significant reduction in levels of alanine transaminase (ALT) (SMD = -26.23; 95% CI = -34.70 to -17.76) and aspartate aminotransferase (AST) (SMD = -19.62; 95% CI = -28.33 to -10.92). Furthermore, reductions in HOMA-IR, body mass index,

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cholesterol, insulin, and waist circumference were observed. The included studies exhibited low to moderate heterogeneity for most outcomes, indicating consistent results across trials. Orlistat significantly improved AST, ALT, and some other metabolic parameters in MASLD patients, suggesting its potential as an additional treatment option. However, the outcome must be interpreted cautiously, considering study heterogeneity. Further high-quality, multicenter research is necessary to confirm these results.

**Keywords:** Metabolic dysfunction-associated fatty liver disease (MA-FLD); metabolic dysfunction-associated steatotic liver disease (MASLD); non-alcoholic fatty liver disease; orlistat; steatosis.

#### Introduction

Nonalcoholic fatty liver disease (NAFLD) involves a range of chronic hepatic conditions closely associated with various extra-hepatic complications, like cardiovascular disorders and metabolic disturbances. NAFLD is primarily described by the aggregation of fat within the liver, accompanied by fibrosis and inflammation. This condition is prone to progress into more severe hepatic diseases, like cirrhosis and liver cancer.<sup>[1,2]</sup> Recent progressions have culminated in the utilization of the designation metabolic dysfunction-associated steatotic liver disease (MASLD), which more accurately encapsulates the fundamental metabolic risk determinants contributing to NAFLD. MASLD is delineated by the existence of hepatic steatosis accompanied by one or more metabolic risk factors, encompassing obesity, type 2 diabetes, or dyslipidemia.<sup>[3,4]</sup> Considering the convergence, these terminologies may be employed interchangeably within populations characterized by metabolic dysfunction, such as individuals with obesity.<sup>[5]</sup> The global prevalence of MASLD is estimated at 32.4%,<sup>[6]</sup> paralleling rising obesity rates and associated metabolic complications.<sup>[7-9]</sup> Although obesity is heterogeneous regarding its etiology, excessive dietary fat consumption is recognized as a significant causative factor in the obesity epidemic.<sup>[10]</sup> Thus, it seems logical to target dietary fat to reduce body mass index (BMI) and decrease the risk of liver steatosis.

Current management for MASLD is centered on lifestyle modifications and gradual weight loss, causing a decrease in hepatic steatosis and serum liver enzyme concentrations, coupled with improved hepatic inflammation and fibrosis.<sup>[11]</sup> Although lifestyle-induced weight loss patterns, such as dietary restriction, seem to be sufficient in some patients, multiple drugs have been evaluated in MASLD patients.<sup>[12,13]</sup>

Orlistat, a semi-synthetic lipase enzyme inhibitor, is designed to treat obesity by inhibiting the breakdown of long-chain triglycerides, thereby decreasing subsequent absorption of dietary triglycerides in the gastrointestinal lümen.<sup>[14,15]</sup> In 2018, a meta-analysis encompassing seven trials comprising 330 patients with MASLD revealed that orlistat treatment could decrease BMI and enhance biochemical indicators of liver damage.<sup>[16]</sup> In recent years, more published trials on orlistat efficacy in MASLD have become available. This study aimed to conduct a thorough analysis of the effectiveness and safety of using the medication orlistat to treat MASLD.

# **Materials and Methods**

#### Systematic Review Protocol

The protocol for this study was conducted based on the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) guidelines.<sup>[17]</sup> The study protocol has been registered in the PROSPE-RO database (registration ID: CRD42024610117).

## Search Strategy

The search was mainly administered on Scopus, PubMed, Web of Science, and Embase utilizing the following keywords and Boolean operators: "Orlistat," "Xenical," "nonalcoholic fatty liver," "non-alcoholic fatty liver," "non-alcoholic," "fatty," "non-alcoholic fatty liver disease," "nonalcoholic," "fatty liver," and "NAFLD" through May 2024. Additionally, MeSH terms were used for "non-alcoholic fatty liver disease" and "fatty liver" as the primary subject for meta-analysis. A manual search was also administered on Google Scholar to increase sensitivity and to find gray literature. The search was updated in January 2025 using the aforementioned keywords (Fig. 1).

# **Eligibility Criteria**

The titles and abstracts of all articles were reviewed independently by two researchers, without considering time and location limitations. After eliminating duplicate articles, full-text articles were reviewed to identify studies meeting our inclusion criteria, and data from these studies were extracted. The inclusion criteria were studies involving adult patients diagnosed with MASLD based on imaging, biopsy, or established clinical criteria, and randomized or non-randomized controlled trials evaluating the efficacy of Orlistat as a monotherapy. The exclusion criteria were cellular or animal studies, studies using Orlistat in combination with other pharmacological or lifestyle interventions, articles without available full text, cross-sectional studies, editorial letters, case reports, single-arm trials, and systematic reviews. The intervention of interest was Orlistat, administered at standard therapeutic doses, typically 120 mg three times daily, as recommended for managing obesity-related conditions. Comparators included placebo, standard care, or other interventions for MASLD management.

## PICO Strategy

**Population (P):** Patients diagnosed with MASLD/metabolic-associated steatohepatitis (MASH) based on imaging, biopsy, or clinical criteria.

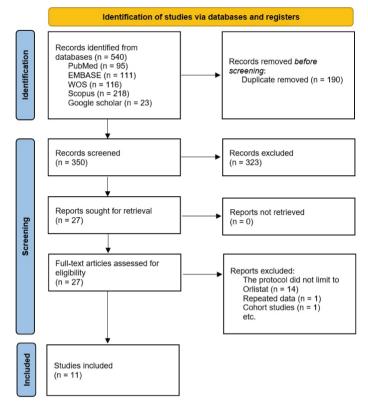


Figure 1. PRISMA flow chart of the screening process.

**Intervention (I):** Orlistat monotherapy, administered at standard therapeutic doses

**Comparator (C):** Placebo, standard care, or other MASLD management interventions.

**Outcome (O):** Changes in liver enzymes (ALT, AST), liver fat content (imaging or biopsy-confirmed), weight loss, and any reported adverse effects.

# **Data Extraction and Bias Assessment**

Two researchers (A.A. and S.K.) independently conducted data extraction (Table 1) and bias assessment (Table S1, Table S2). Extracted information from the studies encompassed general data (first author, year of publication, nation), study characteristics (study design, sample size, follow-up duration), patient characteristics (age, gender, anthropometric indices, insulin levels, glucose, lipid profile, liver function test, diagnostic method), intervention (dosage, duration, dietary regimen), and control (placebo or other interventions). A third investigator (M.R.) resolved any disagreements. Two researchers (A.B. and A.S.) conducted a quality assessment of the included studies independently, and a third researcher (M.R.) resolved and accorded it. Two critical appraisal tools<sup>[18]</sup>—(i) Joanna Briggs Institute (JBI) Prevalence Critical Appraisal Tool and (ii) JBI Critical Appraisal Checklist for Randomized Control/Pseudo-Randomized Trials—were allocated to evaluate the quality of the studies based on their study design. These tools were primarily developed for systematic reviews.

#### **Statistical Analysis**

Data analysis was done using Stata Ver. 14 software. The standardized mean difference (SMD) was used as the effect size, and a 95% confidence interval (CI) was calculated for the combined results. Hetero-

Author, year	Country	Orlistat N	Placebo N	Females N (%)	Mean age, years	Intervention dose	Comparator	Diagnosis	Duration
Randomized clinical trial									
Zahmatkesh et al.[20]	Iran	27	26	23 (43)	13.76	120 mg daily	Placebo	MASLD	3 months
Feng et al. <sup>[21]</sup>	China	40	39	23 (29)	47	120 mg TDS	60 g/day of standard meal replacement powder + reduction of 500–1000 kcal/ day only by controlling their normal diet	MASLD	6 months
Wasta Esmail et al.[22]	Iraq	25	25	30 (60)	43.3	120 mg daily	Placebo	MASLD	3 months
Ye et al. <sup>[23]</sup>	China	68	102	54 (32)	45.58	120 mg TDS	Lifestyle intervention in combination with drug therapy	MASLD	6 months
Harte et al.[24]	UK	8	6	8 (57)	50	120 mg BD	Placebo	MASLD	12 months
Harrison et al.[25]	USA	23	18	28 (68.3)	47	120 mg TDS	Diet/Vitamin E	MASLD	9 months
Zelber-Sagi et al.[26]	Israel	21	23	25 (56)	47.7	120 mg TDS	Placebo	MASLD	6 months
Single arm trial									
Iranparvar Alamdari et al.[27]	Iran	45		24 (53.3)	45.8	120 mg TDS		MASLD	3 months
Khazal et al.[28]	Iraq	60		40 (66.6)	48.5	120 mg TDS		MASLD	4 months
Hussein et al.[29]	Israel	16		10 (62)	42	120 mg TDS		MASH	6 months
Harrison et al.[30]	USA	10		6 (60)	54.4	120 mg TDS		MASLD	6 months

Table 1. Characteristics of included studies

geneity was assessed statistically using Cochran's Q test and I<sup>2</sup> tests. If heterogeneity was present, the appropriate model would be a random-effects model. Heterogeneity intensity was estimated through the I<sup>2</sup> index. If I<sup>2</sup> was less than 25%, the heterogeneity was considered mild; between 25% and 50%, the heterogeneity was moderate; and between 50% and 75%, the heterogeneity was severe. Finally, if the heterogeneity exceeds 75%, it is considered very severe.<sup>[19]</sup> A statistically significant threshold was considered as a p-value<0.05. Besides subgroup analysis and meta-regression, Egger's regression test was also performed to assess publication bias. Finally, sensitivity analyses were conducted to investigate the magnitude of change in the overall effect size and heterogeneity.<sup>[20]</sup>

# Results

# **Trial Characteristics**

The trial characteristics of the 11 included studies can be found in Table 1. Briefly, seven randomized clinical trials were performed between 2004 and 2023 and encompassed 343 patients in the Orlistat group against 239 participants in the control group of RCTs. Single-arm trials encompassed 141 patients in the Orlistat group. The average age of the subjects varied

from 13.76 to 47.7 years, and a majority of them had a 100 percent obesity rate. The timeframe of Orlistat usage in the analyzed studies varied from 3 to 12 months. The studies were conducted across six different countries.

### **Quality of Included Studies**

According to JBI critical appraisal checklists for experimental studies,<sup>[18]</sup> seven included RCTs (Table S1) and four included single-arm trials (Table S2) had a low risk of bias.

## **Quantitative Data Synthesis**

Alanine transaminase (ALT) levels were reported in six trials and four single-arm trials. Orlistat decreased ALT levels in comparison to the initial measurements significantly, both in an analysis of RCTs (SMD=-6.79; 95% CI=-11.96 to -1.63; I<sup>2</sup>=0%, p-value=0.66) and RCTs summed with single-arm trials (SMD=-26.23; 95% CI=-34.70 to -17.76; I<sup>2</sup>=97%, p-value = 0.66) (Fig. S1 and Fig. 2). The funnel plot was proven to be symmetrical, and therefore, no publication bias was present (p-value=0.2067) (Fig. S2).

Due to high heterogeneity in the second analysis, a leave-one-out sensitivity analysis and subgroup analysis were performed based on the

Study	proportion	MRAW 95%-CI
ZELBER–SAGI, 2006 Harrison, 2009 Harte, 2010 Ye, 2019 Feng, 2023 Zahmatkesh, 2023 Harrison, 2004 Hussein, 2007 Khazal, 2007 Iranparvar Alamdari, 2020		-30.60 [-55.83; -5.37] -55.00 [-82.30; -27.70] -27.00 [-44.05; -9.95] -16.50 [-27.12; -5.88] -17.20 [-29.92; -4.48] -9.29 [-11.72; -6.86] -39.00 [-53.51; -24.49] -41.00 [-44.51; -37.49] -14.00 [-17.51; -10.49] -31.60 [-35.60; -27.60]
<b>Random effects model</b> Heterogeneity: $I^2$ = 97%, $\tau^2$ = 145.9210, $p$ < 0.01	-60 -50 -40 -30 -20 -10	- <b>26.23 [-34.70; -17.76]</b> 0

Figure 2. Forest plot describing the connection between ALT concentration in MASLD patients after orlistat administration.

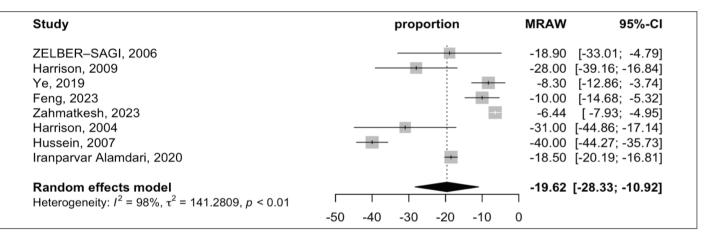


Figure 3. Forest plot describing the connection between AST concentration in MASLD patients after orlistat administration.

age of participants, daily dose of Orlistat, MASLD diagnosis method, and study duration. No study was removed after the sensitivity analysis. ALT changes in both pediatric and adult population studies were significant (Fig. S3), as were all subgroups based on daily doses of Orlistat (Fig. S4). Based on study duration, ALT decreased significantly in subgroups of 4, 6, 9, and 12 months; however, the reduction was not meaningful in the 3-month-long studies (Fig. S5). Finally, for MASLD detection subgroups, biopsy-, MRI–proton density fat fraction (MRI-PDFF)-, and ultrasound-based study results were significant, in contrast with NAFLD activity score (NAS)-based studies (Fig. S6).

Aspartate aminotransferase (AST) levels were reported in five RCTs and three single-arm trials. Orlistat treatment meaningfully reduced AST levels compared to baseline in RCTs (MD=-3.86; 95% CI=-5.64 to -2.08; I<sup>2</sup>=0%, p-value=0.74) and in RCTs summed with single-arm trials (SMD=-19.62; 95% CI=-28.33 to -10.92; I<sup>2</sup>=98%, p-value<0.01) (Fig. S7 and Fig. 3). No publication bias was present due to the symmetrical funnel plot (Fig. S8) (p-value=0.5172).

Due to high heterogeneity in the latter analysis, a leave-one-out sensitivity analysis and subgroup analysis were performed based on the age of participants, daily dose of Orlistat, MASLD diagnosis method, and study duration. Sensitivity analysis did not detect any study as a source of heterogeneity, and no study was removed. All subgroups in the analyses based on age, dose, and study duration were proven significant (Figs. S9, S10, and S11). However, based on the MASLD detection method, AST changes in NAS-based studies were insignificant, in contrast with ultrasound-, MRI-PDFF-, and biopsy-based studies, in which AST changes were considered significant (Fig. S12).

The effect of Orlistat on alkaline phosphatase (ALP) levels was reported in three RCTs. Orlistat did not change ALP levels significantly compared to baseline (SMD=-0.48; 95% CI=-7.05 to 6.10;  $I^2=0\%$ , p-value=0.409) (Fig. 4). Publication bias was not considered significant (p-value = 0.2640) (Fig. S13).

The use of Orlistat was associated with an insignificant reduction in fasting blood sugar (FBS) levels relative to baseline values in MASH and MASLD patients (SMD=-0.28; 95% CI=-0.76 to 0.19) (Fig. S14). Between-study heterogeneity was considered high (I<sup>2</sup>=77%, p-value<0.01). Additionally, funnel plot asymmetry was proven insignificant, and therefore, publication bias was not present (p-value=0.0947) (Fig. S14).

To reduce heterogeneity, subgroup analysis was conducted based on the daily dose of Orlistat and study durations. Based on the Orlistat dose, the changes in FBS were significant in all subgroups except for the 120 mg-TDS subgroup (Fig. S15). Furthermore, FBS change was significant in 3- and 12-month-long studies, in contrast with 6- and 9-month-long

Study	Mean Difference	MD 95%-CI
Harrison, 2009 Ye, 2019 Feng, 2023		4.00 [-5.38; 13.38] -4.20 [-15.61; 7.21] -5.40 [-20.39; 9.59]
<b>Random effects model</b> Heterogeneity: $I^2 = 0\%$ , $\tau^2 = 0.4586$ , $p = 0.43$	-20 -10 0 10	- <b>0.48 [ -7.05; 6.10]</b>

Figure 4. Forest plot describing the connection between ALP concentration in MASLD patients after orlistat administration.

studies, in which FBS change was not considered significant (Fig. S16).

Homeostatic Model Assessment of Insulin Resistance (HOMA-IR) levels were assessed in four studies. In comparison to baseline values, the administration of Orlistat resulted in significantly lower insulin resistance (SMD=-0.67; 95% CI=-0.93 to -0.40; I<sup>2</sup>=0%, p-value=0.53; Fig. S17). In addition, the funnel plot was proven symmetrical (p-value=0.9920) (Fig. S17).

Information gathered from four studies was used to assess body mass index (BMI) after Orlistat administration. In comparison to the control group, Orlistat treatment reduced BMI significantly (SMD=-1.18; 95% CI=-1.45 to -0.90; I<sup>2</sup>=0%, p-value=0.71; Fig. S18). The funnel plot was also proven to be symmetrical, and publication bias was therefore absent (p-value=0.7892) (Fig. S18).

The impact of Orlistat administration on triglyceride (TG) levels in six studies was analyzed, and the result was an insignificant change in TG levels due to Orlistat intake (SMD=0.46; 95% CI=-0.64 to 1.56; I<sup>2</sup>=94%, p-value<0.01) (Fig. S19). Sensitivity analysis detected the study by Feng et al.<sup>[21]</sup> as a source of heterogeneity, and after its removal, heterogeneity decreased from 94% to 56%, but changes in TG levels were still insignificant (SMD=-0.07; 95% CI=-0.46 to 0.32; I<sup>2</sup>=56%, p-value=0.06; Fig. S20). The funnel plot was also proven symmetrical (p-value=0.7803) (Fig. S19).

Cholesterol levels were assessed in seven studies before and after Orlistat intake. Our analysis showed that Orlistat did not change cholesterol levels significantly (SMD=-0.11; 95% CI=-0.68 to 0.46; I<sup>2</sup>=86%, p-value<0.01) (Fig. S21). However, after performing a sensitivity analysis, the study by Feng et al.<sup>[21]</sup> was removed from this analysis, and alongside a significant reduction in heterogeneity, the results changed (SMD=-0.38; 95% CI=-0.67 to -0.08; I<sup>2</sup>=34%, p-value=0.18) (Fig. S22). The funnel plot was also proven symmetrical (p-value=0.9139) (Fig. S21).

Analysis of insulin levels in five studies showed that Orlistat treatment caused a slight decrease in insulin levels in MASLD or MASH patients (SMD=-2.33; 95% CI=-3.44 to -1.21; I<sup>2</sup>=0%, p-value=0.76) (Fig. S23). Funnel plot asymmetry was also proven insignificant (p-value=0.9881) (Fig. S23).

The effect of Orlistat on low-density lipoprotein (LDL) levels was evaluated in six studies. Our analysis demonstrated that Orlistat administration did not significantly influence LDL levels (SMD=-0.55; 95% CI=-1.14 to 0.04; I<sup>2</sup>=86%, p-value<0.01; Fig. S24). Sensitivity analysis repeatedly found the study by Feng et al.<sup>[21]</sup> as a source of heterogeneity, and after its removal, the heterogeneity reached 42%; however, the reduction was still insignificant (SMD=-0.24; 95% CI=-0.58 to 0.11; I<sup>2</sup>=42%, p-value=0.14; Fig. S25). As a result of funnel plot symmetry, publication bias was absent (p-value=0.9139) (Fig. S24).

Table 2. Meta-regression results				
Comparison	Age	Gender	BMI	
ALT	0.1154	0.0491	0.0559	
AST	0.2312	<0.0001	0.1067	
ALP	0.4705	0.1928	0.1925	
FBS	0.3522	0.6726	0.5614	
HOMA-IR	0.2629	0.6621	0.4055	
BMI	0.4791	0.7841	0.3281	
TG	0.2455	0.2484	0.8192	
Cholesterol	0.2455	0.2484	0.8192	
Insulin	0.5421	0.5208	0.9901	
LDL	0.9146	0.3520	0.4505	
WC	0.7541	0.8350	0.6671	
SBP	0.6967	0.5432	0.5581	

BMI: Body mass index; ALT: Alanine transaminase; AST: Aspartate aminotransferase; ALP: Alkaline phosphatase; FBS: Fasting blood sugar; HOMA-IR: Homeostatic Model Assessment of Insulin Resistance; TG: Triglyceride; LDL: Low-density lipoprotein; WC: Waist circumference; SBP: Systolic blood pressure.

The effect of Orlistat on waist circumference (WC) in five studies was analyzed. Finally, waist circumference was reduced compared to baseline after Orlistat treatment (SMD=-3.08; 95% CI=-3.94 to -2.22; I<sup>2</sup>=0%, p-value=0.94) (Fig. S26). The funnel plot was also considered symmetrical (p-value=0.5382) (Fig. S26).

The effect of daily Orlistat intake on systolic blood pressure (SBP) was analyzed in three studies. Orlistat treatment did not yield a statistically significant alteration in SBP compared to the initial measurement (SMD=2.34; 95% CI=-1.11 to 5.79; I<sup>2</sup>=0%, p-value=0.80) (Fig. S27). Publication bias was absent in this analysis (p-value=0.9139) (Fig. S27).

As sensitivity analyses were performed to assess the robustness of results, data are available within the supplementary material (Tables S3–S8).

Meta-regression analysis was also conducted to explore the potential influence of age, BMI, and sex on the 12 previously specified comparisons by assessing p-values for them.

In more detail, gender structure showed a significant correlation with changes in ALT and AST (Table 2). The rest of the comparisons were insignificant.

# Discussion

Orlistat exerts an inhibitory effect on pancreatic and gastric lipases by forming a covalent link with serine residues within their active sites. [14,32-33] As these lipases act as major components of dietary lipid catabolism, inhibiting them will diminish the amount of available free fatty acids in the liver; this will also result in controlled blood pressure.[31,33] Moreover, a recent meta-analysis of seven trials demonstrated that Orlistat administration is beneficial in maintaining uric acid at a low level. <sup>[34]</sup> Although these findings suggest that Orlistat may be advantageous for MASLD patients' hepatic status, Orlistat usage was also correlated with rare and scattered complications, including several hepatic diseases (e.g., cholestatic hepatitis and subacute liver failure).<sup>[32]</sup> To gain a better understanding of Orlistat's efficacy, this systematic review and meta-analysis summarized its effects on MASLD and MASH patients. This study encompassed eleven RCTs and single-arm trials with a total of 592 participants. Quantitative synthesis results from the included studies showed that Orlistat treatment significantly improved metabolic and biochemical parameters of liver disease, including ALT, AST, BMI, cholesterol, HOMA-IR, and WC, while other anthropometric and metabolic parameters showed non-significant changes.

Dyslipidemia is recognized as a prominent risk factor for MASLD.<sup>[35]</sup> Improvement of dyslipidemia is a priority in the supervision of MASLD patients.<sup>[36]</sup> A meta-analysis of 33 RCTs confirmed a reduction in total cholesterol and TG, which was consistent with our findings.[37] Using meta-regression analysis, they further reported that the lipid-lowering effect of Orlistat may be caused by its concomitant effect on body weight loss.<sup>[37]</sup> In non-obese subjects, the baseline triglyceride level is a strong indicator of the development and regression of MASLD.[38] Orlistat is a lipase inhibitor that inhibits the hydrolysis of triglycerides and leads to lower fatty acid absorption.<sup>[39]</sup> Its effects are not limited to liver enzymes and parameters assessed in this analysis; liver fat content, NAS, and NAFLD fibrosis score (NFS) are some other vital scores in evaluating MASLD severity. Orlistat failed to improve NAS and NFS scores in the studies by Harrison et al.,<sup>[25]</sup> Esmail et al.,<sup>[40]</sup> and Ali Khan et al.<sup>[41]</sup> In contrast, the ultrasound properties of the liver improved after Orlistat administration in the studies by Khazal et al.<sup>[28]</sup> and Ali Khan et al.<sup>[41]</sup> Liver fat content showed paradoxical results after Orlistat consumption; while in the study by Feng et al.,<sup>[21]</sup> MRI-PDFF analysis showed significant improvement in the Orlistat group, a whole-liver analysis based on MRI-PDFF in the study by Ye et al.<sup>[23]</sup> demonstrated that fat content was not altered in the Orlistat group compared to the control group.

Although lifestyle modifications and weight loss are important elements in the management of MASLD, recent advances in pharmacological therapies are reshaping the treatment landscape. Resmetirom, a thyroid hormone receptor- $\beta$  agonist, has received conditional approval for patients with MASLD and significant fibrosis (F2-F3), showing promise in reducing liver fat and improving histology.<sup>[42]</sup> Similarly, semaglutide, a GLP-1 receptor agonist, has demonstrated significant weight loss and liver histological improvements, particularly in the ESSENCE-3 trial. <sup>[43]</sup> These innovations highlight the growing role of pharmacotherapy alongside lifestyle interventions, emphasizing the need for a tailored, multidisciplinary approach to MASLD treatment. In the current study, we focused on the effect of Orlistat on several metabolic parameters. Our analysis demonstrated its significant reduction of ALT, which may be a proper therapeutic pathway, as a reduction of ALT has been shown to be correlated with significant histological improvement in liver biopsies.<sup>[44]</sup> Additionally, our analysis showed that Orlistat reduced AST levels in MASLD patients. It has been claimed that a reduction in AST

is associated with improved liver function and a lower risk of fibrosis and cirrhosis.<sup>[45,46]</sup> Other improved parameters, like BMI and cholesterol, are also correlated with a lower risk of comorbid conditions such as obesity or metabolic syndrome.<sup>[44]</sup>

MASLD and metabolic syndrome are strongly connected, with MASLD having a bidirectional relationship with metabolic syndrome.<sup>[47,48]</sup> Type 2 diabetes, as a main risk factor for MASLD, is correlated with the progression of fibrosis and disease deterioration. It is also associated with liver-related mortality.<sup>[49]</sup> According to current international clinical practice guidelines, the assessment of insulin resistance is considered a crucial component during the diagnosis and prognosis determination of individuals with MASLD, along with numerous other factors.[50-52] Insulin resistance and subsequent hyperinsulinemia increase free fatty acid delivery and fat accumulation, and stimulate anabolic metabolism. <sup>[53]</sup> Recently, multiple meta-analyses evaluated the impact of Orlistat on metabolic parameters. Data from Zhou et al.[54] showed that Orlistat treatment is paralleled by a significant reduction in FBS. A meta-analysis on obese and overweight subjects reported that Orlistat significantly improved hemoglobin A1C (HbA1c) and FBS, compared to placebo. <sup>[55]</sup> Our results further showed improvement in HOMA-IR following Orlistat treatment. Moreover, FBS and insulin levels were significantly improved following the administration of Orlistat.

Despite the similar design of the included studies, heterogeneity was found among studies. The average age of included participants was inconsistent, with a range of 13.76 to 54.4 years, and to resolve this issue, we conducted a meta-regression featuring participants' mean age in the encompassed studies. The results showed that age did not affect alterations in the previously mentioned comparisons. The duration of treatment was not uniform and ranged from three to twelve months in the included studies. Subsequently, a duration-based subgroup analysis was conducted to assess the effect of study duration on our analyses. Moreover, the MASLD detection methods encompassed a range of techniques (e.g., MRI-PDFF, NAS, ultrasound), which could affect our results. Due to this issue, subgroup analyses were performed based on the detection method to minimize the effect.

Finally, our study has some advantages, including the following: strict exclusion and inclusion criteria were used, and only clinical trials were included. We conducted a systematic and comprehensive literature search to reduce bias. However, this study had some limitations, which might affect the generalizability of the study results. The number of studies included in this systematic review and meta-analysis was limited, and a relatively large sample size was not observed within them. These limitations may influence our meta-analysis results. To acquire more robust evidence to guide clinical practice, additional high-quality and multicenter research should be conducted.

#### Conclusion

In this study, Orlistat's potential for MASLD management was quantitatively analyzed. Orlistat significantly improved AST levels and some metabolic parameters in MASLD patients. It was demonstrated that Orlistat might provide an additional treatment choice for patients with MASLD. These findings need to be interpreted with caution.

#### 

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Author Contributions: Concept – MR; Design – MR; Supervision – SN, MR; Data Collection and/or Processing – IEV, AB, SK, AS, MRaz, YR; Analysis and/ or Interpretation – IEV, MN, MAB, MMMN; Literature Search – MR, AA; Writing – AA, YR, MRaz; Critical Reviews – MR, AA.

Peer-review: Externally peer-reviewed.

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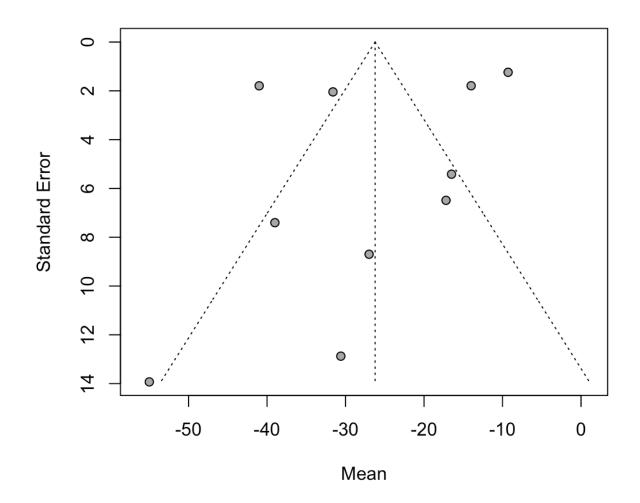
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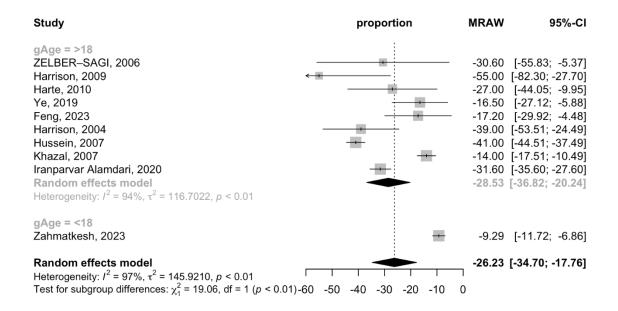
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Study	Mean Difference	MD 95%-CI
ZELBER–SAGI, 2006 Harrison, 2009 Harte, 2010 Ye, 2019 Feng, 2023		-17.90 [-45.38; 9.58] -10.00 [-41.97; 21.97] - 2.00 [-44.67; 48.67] -10.90 [-22.45; 0.65] -12.20 [-26.15; 1.75]
Zahmatkesh, 2023		-4.26 [-7.24; -1.28]
<b>Random effects model</b> Heterogeneity: $I^2 = 0\%$ , $\tau^2 = 8.3522$ , $p = 0.66$	-40 -20 0 20 40	-6.79 [-11.96; -1.63]

Figure S1. Forest plot describing the effect of orlistat administration on ALT levels (in RCTs)



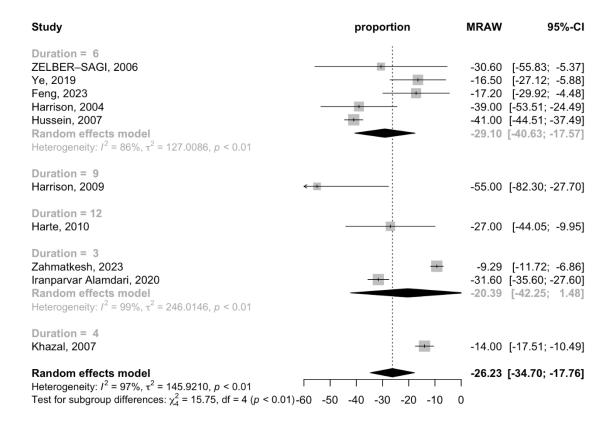
**Figure S2.** Funnel plot for the studies describing the effect of orlistat administration on ALT levels (in RCTs and single-arm trials)



**Figure S3.** Forest plot describing the effect of orlistat administration on ALT levels (in RCTs and singlearm trials) (Subgroup analysis based on age of participants)

Study	proportion	MRAW	95%-CI
Dose = 120 mg TDS ZELBER–SAGI, 2006 Harrison, 2009 Ye, 2019 Feng, 2023 Harrison, 2004 Hussein, 2007 Khazal, 2007 Random effects model Heterogeneity: $l^2$ = 95%, $\tau^2$ = 166.4146, <i>p</i> < 0.01		-55.00 [-6 -16.50 [- -17.20 [- -39.00 [-5 -41.00 [-4 -14.00 [-1	55.83; -5.37] 32.30; -27.70] 27.12; -5.88] 29.92; -4.48] 53.51; -24.49] 14.51; -37.49] 17.51; -10.49] 9.46; -17.54]
Dose = 120 mg BD Harte, 2010 Iranparvar Alamdari, 2020 Random effects model Heterogeneity: $I^2 = 0\%$ , $\tau^2 = 0$ , $p = 0.61$	*	-31.60 [-3	44.05; -9.95] 35.60; -27.60] 5.26; -27.46]
<b>Dose = 120 mg daily</b> Zahmatkesh, 2023 <b>Random effects model</b> Heterogeneity: $l^2 = 97\%$ , $\tau^2 = 145.9210$ , $p < 0.01$ Test for subgroup differences: $\chi^2_2 = 93.99$ , df = 2 ( $p < 0.01$ ).	-60 -50 -40 -30 -20 -10		11.72; -6.86] • <b>4.70; -17.76]</b>

**Figure S4.** Forest plot describing the effect of orlistat administration on ALT levels (in RCTs and singlearm trials) (Subgroup analysis based on intake dose)

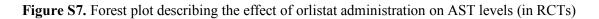


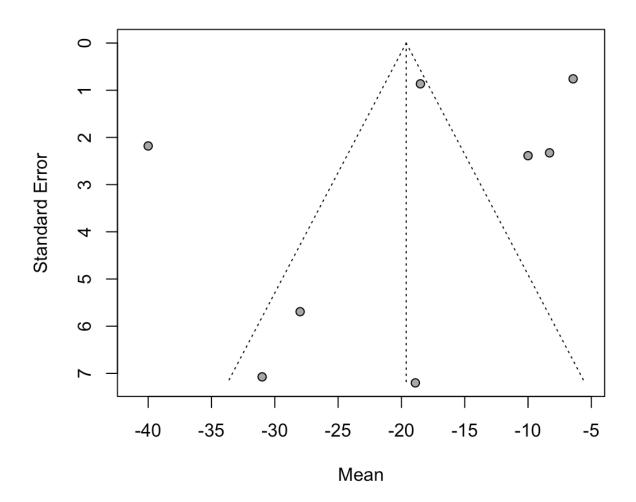
**Figure S5.** Forest plot describing the effect of orlistat administration on ALT levels (in RCTs and singlearm trials) (Subgroup analysis based on study duration)

Study	proportion	MRAW	95%-CI
Assessment = biopsy ZELBER–SAGI, 2006 Harte, 2010 Harrison, 2004 Hussein, 2007 Random effects model Heterogeneity: $J^2$ = 3%, $\tau^2$ = 10.0163, $p$ = 0.38		-27.00 [-4 -39.00 [-5 -41.00 [-4	55.83; -5.37] 14.05; -9.95] 3.51; -24.49] 4.51; -37.49] 4.62; -32.71]
Assessment = NAS Harrison, 2009 Zahmatkesh, 2023 Random effects model Heterogeneity: $J^2$ = 91%, $\tau^2$ = 946.9288, $p$ < 0.01		-9.29 [-	2.30; -27.70] 11.72; -6.86] 4.64; 14.56]
Assessment = MRI-PDFF Ye, 2019 Feng, 2023 Random effects model Heterogeneity: $J^2 = 0\%$ , $\tau^2 = 0$ , $p = 0.93$		-17.20 [-2	27.12; -5.88] 29.92; -4.48] 24.94; -8.64]
Assessment = US Khazal, 2007 Iranparvar Alamdari, 2020 Random effects model Heterogeneity: $l^2$ = 98%, $\tau^2$ = 151.1914, $\rho$ < 0.01	*	-31.60 [-3	7.51; -10.49] 5.60; -27.60] 0.02; -5.52]
<b>Random effects model</b> Heterogeneity: $l^2 = 97\%$ , $\tau^2 = 145.9210$ , $p < 0.01$ Test for subgroup differences: $\chi_3^2 = 18.87$ , df = 3 ( $p < 0.01$ )-6	60 -50 -40 -30 -20 -10	<b>-26.23 [-3</b>	4.70; -17.76]

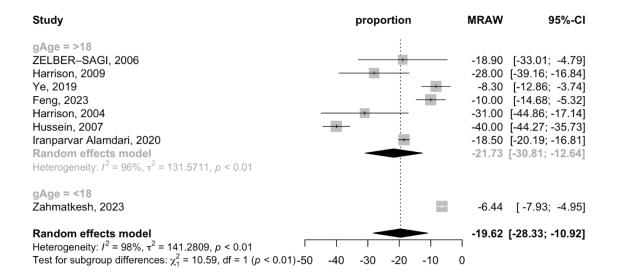
**Figure S6.** Forest plot describing the effect of orlistat administration on ALT levels (in RCTs and singlearm trials) (Subgroup analysis based on NAFLD detection method)

Study	Mean Difference	MD 95%-CI
ZELBER–SAGI, 2006 Harrison, 2009 Ye, 2019 Feng, 2023		-10.10 [-25.87; 5.67] 1.00 [-14.73; 16.73] -5.10 [-9.99; -0.21] -5.90 [-11.48; -0.32]
Zahmatkesh, 2023	-	-3.33 [-5.41; -1.25]
<b>Random effects model</b> Heterogeneity: $l^2 = 0\%$ , $\tau^2 = 0$ , $p = 0.74$	-20 -10 0 10 20	-3.86 [-5.64; -2.08]

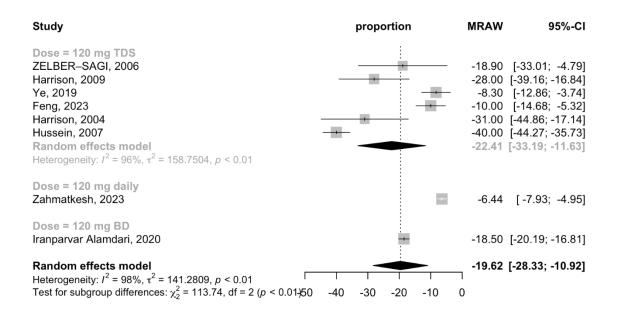




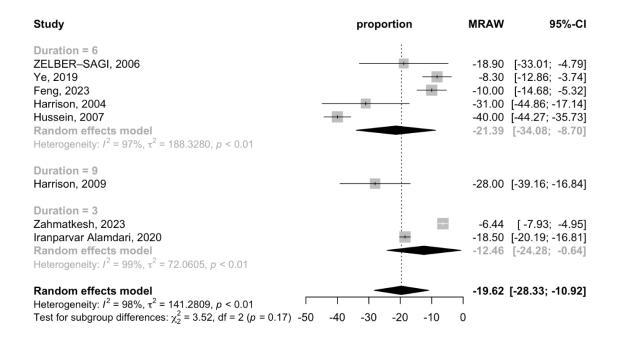
**Figure S8.** Funnel plot for the studies describing the effect of orlistat administration on AST levels (in RCTs and single-arm trials)



**Figure S9.** Forest plot describing the effect of orlistat administration on AST levels (in RCTs and singlearm trials) (Subgroup analysis based on age of participants)



**Figure S10.** Forest plot describing the effect of orlistat administration on AST levels (in RCTs and single-arm trials) (Subgroup analysis based on intake dose)



**Figure S11.** Forest plot describing the effect of orlistat administration on AST levels (in RCTs and single-arm trials) (Subgroup analysis based on study duration)

Study	proportion	MRAW	95%-CI
Assessment = biopsy ZELBER–SAGI, 2006 Harrison, 2004 Hussein, 2007 Random effects model Heterogeneity: $J^2$ = 77%, $\tau^2$ = 88.9152, $p$ = 0.01	*	-31.00 [-4 -40.00 [-4	33.01; -4.79] 4.86; -17.14] 4.27; -35.73] 3.81; -19.00]
Assessment = NAS Harrison, 2009 Zahmatkesh, 2023 Random effects model Heterogeneity: $J^2$ = 93%, $\tau^2$ = 215.9274, $p$ < 0.01	-	-6.44 [·	9.16; -16.84] -7.93; -4.95] 7.56; 4.60]
Assessment = MRI-PDFF Ye, 2019 Feng, 2023 Random effects model Heterogeneity: $l^2 = 0\%$ , $\tau^2 = 0$ , $p = 0.61$		-10.00 [-1	2.86; -3.74] 4.68; -5.32] 2.39; -5.86]
Assessment = US Iranparvar Alamdari, 2020	+	-18.50 [-2	0.19; -16.81]
Random effects model           Heterogeneity: $l^2$ = 98%, $\tau^2$ = 141.2809, $p < 0.01$ Test for subgroup differences: $\chi_3^2$ = 30.39, df = 3 ( $p < 0.01$ )-5	0 -40 -30 -20 -10	<b>-19.62 [-2</b> 8	3.33; -10.92]

**Figure S12.** Forest plot describing the effect of orlistat administration on AST levels (in RCTs and single-arm trials) (Subgroup analysis based on NAFLD detection method)

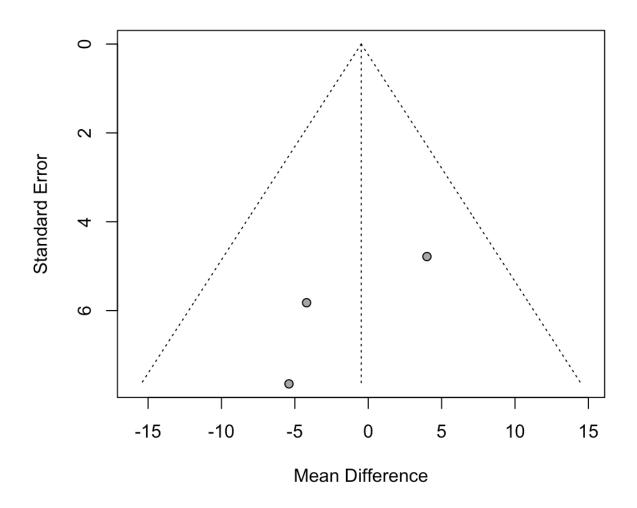


Figure S13. Funnel plot for the studies describing the effect of orlistat administration on ALP levels

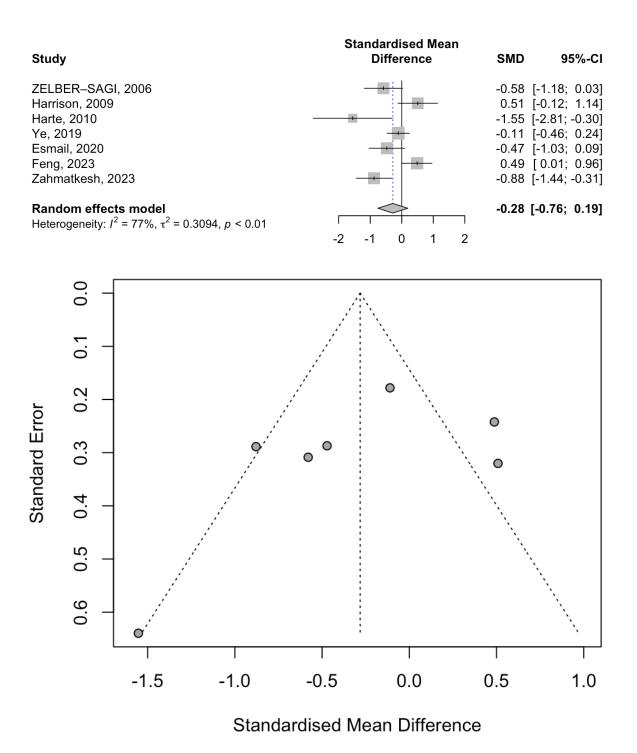


Figure S14. Forest and funnel plot for the studies describing the effect of orlistat administration on FBS

Study	Standardised Mean Difference	SMD	95%-CI
Dose = 120 mg TDS ZELBER–SAGI, 2006		-0 58	[-1.18; 0.03]
Harrison, 2009			[-0.12; 1.14]
Ye, 2019			[-0.46; 0.24]
Feng, 2023			[ 0.01; 0.96]
Random effects model		0.08 [	[-0.41; 0.56]
Heterogeneity: $I^2 = 71\%$ , $\tau^2 = 0.1745$ , $p = 0.02$			
Dose = 120 mg BD			
Harte, 2010 –		-1.55	[-2.81; -0.30]
Dose = 120 mg daily		0.47	[ 1 02: 0 00]
Esmail, 2020 Zahmatkesh, 2023			[-1.03; 0.09] [-1.44; -0.31]
Random effects model			[ <b>-1.07; -0.27]</b>
Heterogeneity: $l^2 = 0\%$ , $\tau^2 = < 0.0001$ , $p = 0.32$	$\sim$	0.01 [	[
Random effects model		-0.28	[-0.76; 0.19]
Heterogeneity: $I^2 = 77\%$ , $\tau^2 = 0.3094$ , $p < 0.01$			
Test for subgroup differences: $\chi_2^2 = 8.72$ , df = 2 ( $p = 0.01$ )	-2 -1 0 1 2		

**Figure S15.** Forest plot for the studies describing the effect of orlistat administration on FBS (Subgroup analysis based on intake dose)

Study	Standardised Mean Difference	SMD	95%-CI
<b>Duration = 6</b> ZELBER–SAGI, 2006 Ye, 2019 Feng, 2023 <b>Random effects model</b> Heterogeneity: $l^2$ = 75%, $\tau^2$ = 0.1973, $p$ = 0.02		-0.11 [- 0.49 [	1.18; 0.03] 0.46; 0.24] 0.01; 0.96] <b>0.62; 0.53]</b>
<b>Duration = 9</b> Harrison, 2009		0.51 [-	0.12; 1.14]
Duration = 12 Harte, 2010 —		-1.55 [-	2.81; -0.30]
<b>Duration = 3</b> Esmail, 2020 Zahmatkesh, 2023 <b>Random effects model</b> Heterogeneity: $I^2 = 0\%$ , $\tau^2 = < 0.0001$ , $p = 0.32$		-0.88 [-	1.03; 0.09] 1.44; -0.31] <b>1.07; -0.27]</b>
<b>Random effects model</b> Heterogeneity: $l^2 = 77\%$ , $\tau^2 = 0.3094$ , $p < 0.01$ Test for subgroup differences: $\chi_3^2 = 14.30$ , df = 3 ( $p < 0.01$ )	-2 -1 0 1 2	-0.28 [-	0.76; 0.19]

**Figure S16.** Forest plot for the studies describing the effect of orlistat administration on FBS (Subgroup analysis based on study duration)

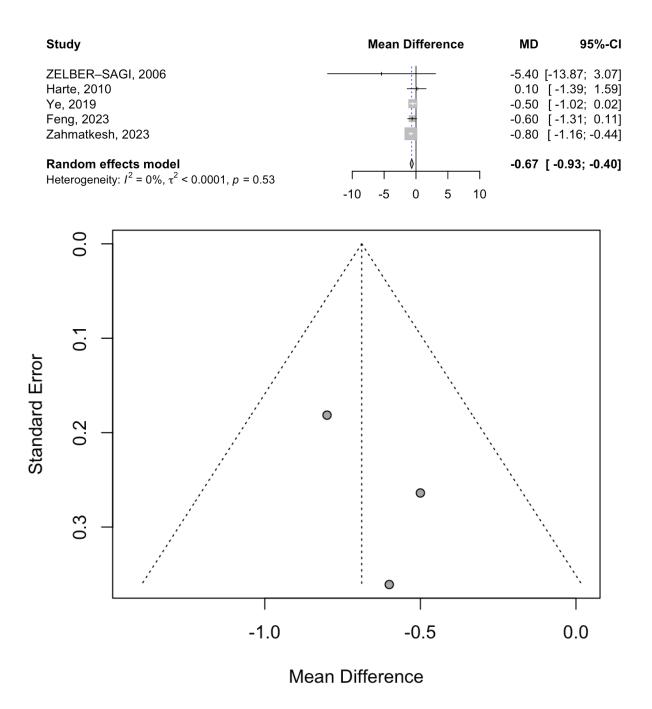


Figure S17. Forest and funnel plot for studies describing the effect of orlistat administration on HOMA

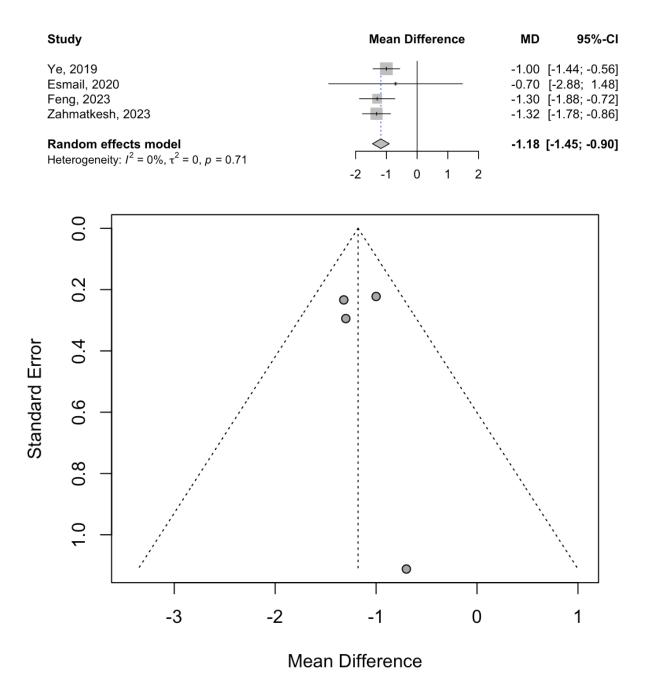
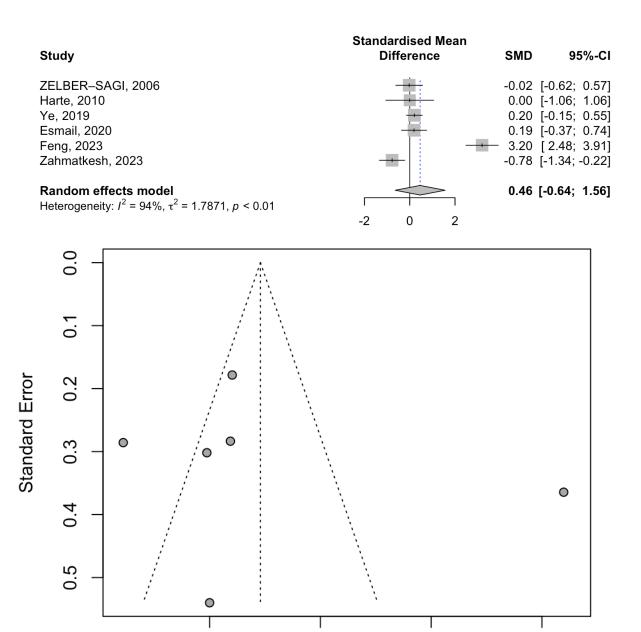


Figure S18. Forest and funnel plot for studies describing the effect of orlistat administration on BMI



Standardised Mean Difference

Figure S19. Forest and funnel plot for studies describing the effect of orlistat administration on TG levels

Study	Standardised Mean Difference	SMD 95%-CI
ZELBER–SAGI, 2006 Harte, 2010 Ye, 2019 Esmail, 2020 Zahmatkesh, 2023		-0.02 [-0.62; 0.57] 0.00 [-1.06; 1.06] 0.20 [-0.15; 0.55] 0.19 [-0.37; 0.74] -0.78 [-1.34; -0.22]
<b>Random effects model</b> Heterogeneity: $I^2$ = 56%, $\tau^2$ = 0.1083, $p$ = 0.06	-1 -0.5 0 0.5 1	-0.07 [-0.46; 0.32]

**Figure S20.** Forest plot for studies describing the effect of orlistat administration on TG levels (after removal of Feng's study)

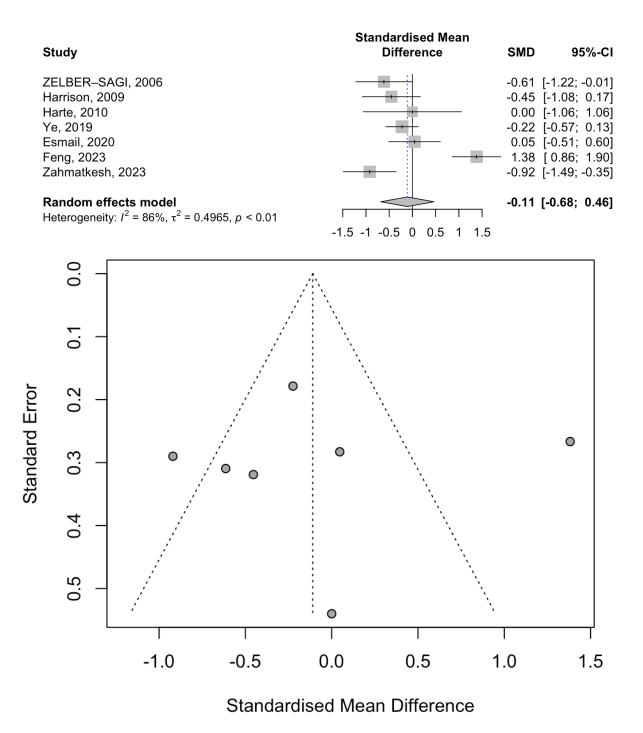


Figure S21. Forest and funnel plot for studies describing the effect of orlistat administration on cholesterol levels

Study	Standardised Mean Difference	SMD	95%-CI
ZELBER–SAGI, 2006 Harrison, 2009 Harte, 2010 Ye, 2019		-	
Esmail, 2020 Zahmatkesh, 2023		-	.51; 0.60]
<b>Random effects model</b> Heterogeneity: $l^2 = 34\%$ , $\tau^2 = 0.0469$ , $p = 0.18$	-1 -0.5 0 0.5 1	-0.38 [-0.	67; -0.08]

**Figure S22.** Forest plot for studies describing the effect of orlistat administration on cholesterol levels (after removal of Feng's study)

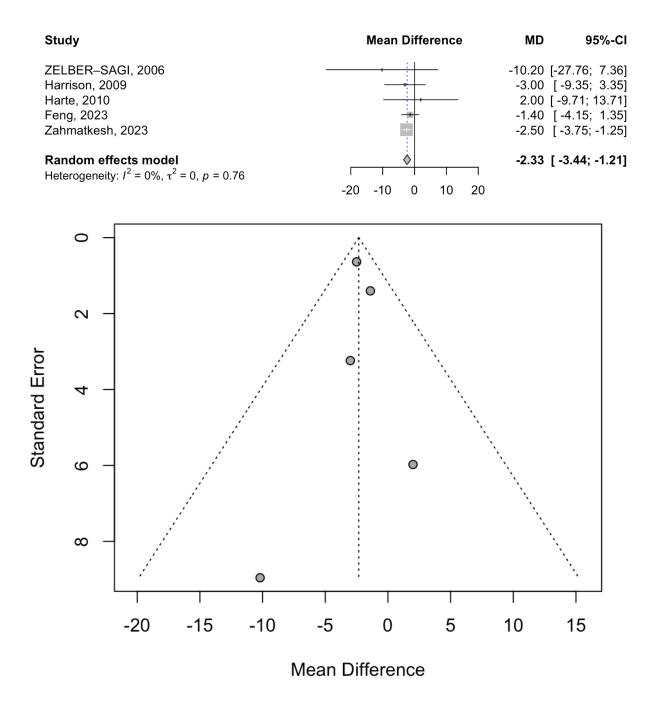
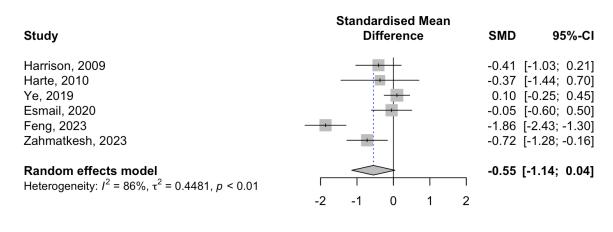
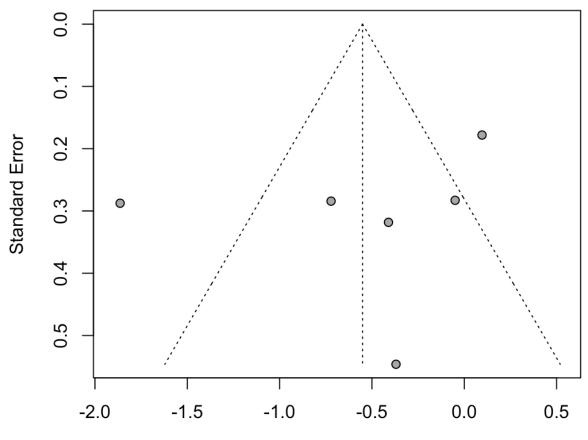


Figure S23. Forest and funnel plot for studies describing the effect of orlistat administration on insulin levels





Standardised Mean Difference Figure S24. Forest and funnel plot for studies describing the effect of orlistat administration on LDL levels

Study	Standardised Mean Difference	SMD	95%-CI
Harrison, 2009		-0.41 [-	1.03; 0.21]
Harte, 2010		-0.37 [-	1.44; 0.70]
Ye, 2019		0.10 [-	0.25; 0.45]
Esmail, 2020		-0.05 [-	0.60; 0.50]
Zahmatkesh, 2023		-0.72 [-	1.28; -0.16]
Random effects model		-0.24 [-0	0.58; 0.11]
Heterogeneity: $I^2 = 42\%$ , $\tau^2 = 0.0667$ , $p = 0.14$			
	-1 -0.5 0 0.5 1		

**Figure S25.** Forest plot for studies describing the effect of orlistat administration on LDL levels (after removal of Feng's study)

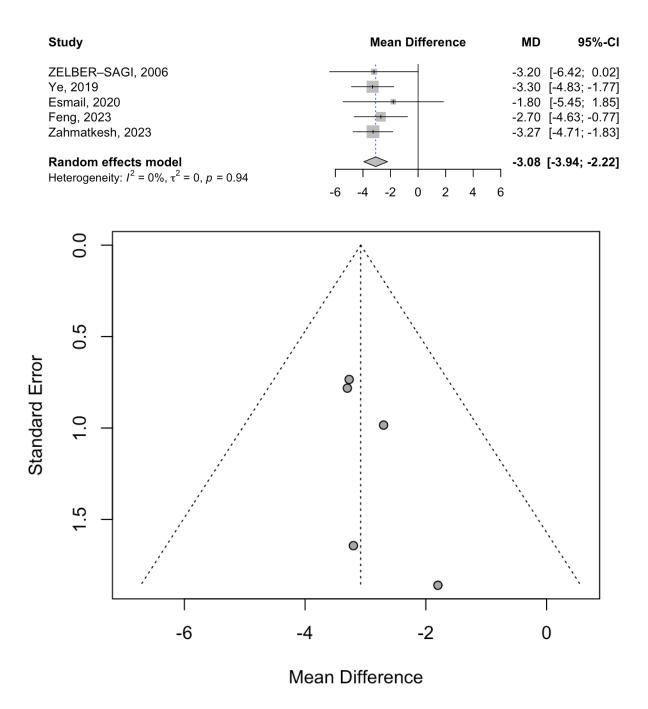


Figure S26. Forest and funnel plot for studies describing the effect of orlistat administration on WC

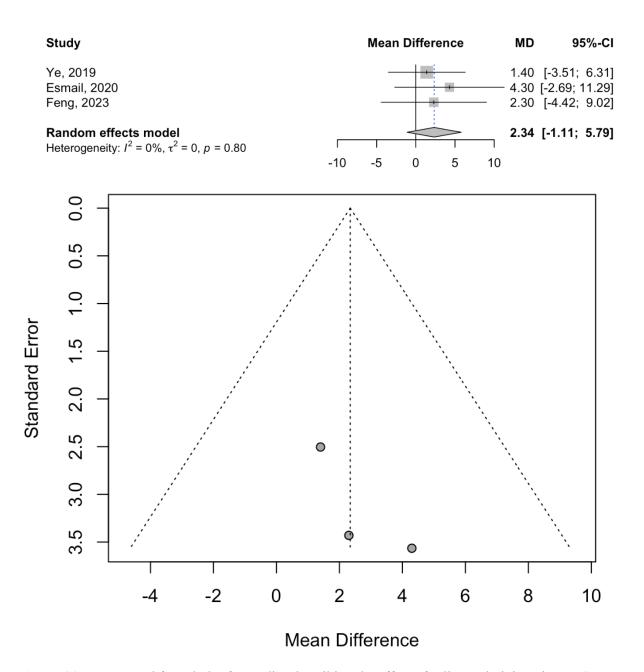


Figure S27. Forest and funnel plot for studies describing the effect of orlistat administration on SBP

RCT studies included	1	2	3	4	5	6	7	8	9	10	11	12	13
Zahmatkesh et al.(15)	Yes	No	Yes	No	No	Yes							
Feng et al.(16)	Yes	Yes	Yes	Yes	Yes	Yes	n/a	Yes	Yes	Yes	Yes	Yes	Yes
Wasta Esmail et al.(14)	Yes	No	Yes	No	No	Yes	n/a	Yes	Yes	Yes	Yes	Yes	Yes
Ye et al. (17)	Yes	No	Yes	No	No	Yes	n/a	Yes	Yes	Yes	Yes	Yes	Yes
Harte et al. (18)	No	Yes	Yes	Yes	Yes	Yes	No	Yes	Yes	Yes	Yes	Yes	Yes
Harrison et al. (19)	Yes	No	Yes	No	No	Yes							
Zelber–Sagi et al.(20)	Yes	Yes	Yes	Yes	Yes	Yes	n/a	Yes	Yes	Yes	Yes	Yes	Yes

Table S1. JBI critical appraisal checklist for randomized controlled trials

- 1- Was true randomization used for assignment of participants to treatment groups?
- 2- Was allocation to treatment groups concealed?
- 3- Were treatment groups similar at the baseline?
- 4- Were participants blind to treatment assignment?
- 5- Were those delivering the treatment blind to treatment assignment?
- 6- Were treatment groups treated identically other than the intervention of interest?
- 7- Were outcome assessors blind to treatment assignment?
- 8- Were outcomes measured in the same way for treatment groups?
- 9- Were outcomes measured in a reliable way?
- 10- Was follow up complete and if not, were differences between groups in terms of their follow up adequately described and analyzed?
- 11- Were participants analyzed in the groups to which they were randomized?
- 12- Was appropriate statistical analysis used?
- 13- Was the trial design appropriate and any deviations from the standard RCT design (individual randomization, parallel groups) accounted for in the conduct and analysis of the trial?

Single-arm trials included	1	2	3	4	5	6	7	8	9
Iranparvar Alamdari et al. (21)	Yes	No	Yes						
Khazal et al.(13)	Yes	No	Yes						
Hussein et al. (22)	Yes	No	Yes						
Harrison et al. (23)	Yes	No	Yes						

Table S2. JBI critical appraisal checklist for quasi-experimental studies

- 1- Is it clear in the study what is the "cause" and what is the "effect" (i.e. there is no confusion about which variable comes first)?
- 2- Was there a control group?
- 3- Were participants included in any comparisons similar?
- 4- Were the participants included in any comparisons receiving similar treatment/care, other than the exposure or intervention of interest?
- 5- Were there multiple measurements of the outcome, both pre and post the intervention/exposure?
- 6- Were the outcomes of participants included in any comparisons measured in the same way?
- 7- Were outcomes measured in a reliable way?
- 8- Was follow-up complete and if not, were differences between groups in terms of their follow-up adequately described and analyzed?
- 9- Was appropriate statistical analysis used?

Study_excluded	Estimate	CI_lb	CI_ub	Tau2	12
Excluding ZELBER-SAGI	-26.0276	-35.055	-17.0002	157.4491	97.0253
Excluding Harrison	-24.5128	-32.7724	-16.2532	128.6438	96.9631
Excluding Harte	-26.2896	-35.5969	-16.9824	164.0992	97.026
Excluding Ye	-27.4836	-36.7498	-18.2173	158.0325	97.0256
Excluding Feng	-27.3356	-36.6087	-18.0626	159.7347	97.029
Excluding Zahmatkesh	-28.5317	-36.8226	-20.2409	116.7022	93.8476
Excluding Harrison	-24.9138	-33.7777	-16.0499	144.684	96.9616
Excluding Hussein	-23.7541	-32.005	-15.5032	115.5954	92.6761
Excluding Khazal	-27.9846	-37.0439	-18.9253	145.9706	96.8344
Excluding Iranparvar Alamdari	-25.6198	-35.1247	-16.115	164.3475	96.6024

Table S3. Sensitivity analysis for proportional ALT analysis

**Table S4.** Sensitivity analysis for proportional AST analysis

Study_Excluded	Estimate	CI_lb	CI_ub	Tau2	I2
Excluding ZELBER-SAGI	-19.7693	-29.5406	-9.99791	161.3766	97.9121
Excluding Harrison	-18.5779	-28.146	-9.00967	152.27	97.8667
Excluding Ye	-21.3846	-30.7796	-11.9897	142.6583	97.8805
Excluding Feng	-21.1369	-30.7341	-11.5397	149.6008	97.9017
Excluding Zahmatkesh	-21.7292	-30.815	-12.6435	131.5711	95.5072
Excluding Harrison	-18.3224	-27.6337	-9.01115	145.3561	97.87
Excluding Hussein	-15.7354	-22.4282	-9.04265	65.50655	95.3582
Excluding Iranparvar Alamdari	-19.8825	-30.0081	-9.7569	167.5124	97.4112

**Table S5.** Sensitivity analysis for studies assessing TG levels

Study	Estimate	CI_lb	CI_ub	p_value	Tau2	I2
ZELBER–SAGI	0.5589	-0.77799	1.895794	0.41257	2.206851	96.29354
Harte	0.544935	-0.77555	1.86542	0.418611	2.187304	96.79489
Ye	0.514243	-0.84422	1.872707	0.458124	2.269922	95.36896
Esmail	0.516276	-0.83592	1.868477	0.454266	2.258144	96.26787
Feng	-0.07074	-0.46142	0.319943	0.722672	0.108736	57.7949
Zahmatkesh	0.712844	-0.50195	1.927641	0.2501	1.799965	95.37402

Study	Estimate	CI_lb	CI_ub	p_value	Tau2	I2
ZELBER-SAGI	-0.02513	-0.672	0.621734	0.939294	0.55273	87.65158
Harrison	-0.05281	-0.71398	0.608357	0.875598	0.582644	88.31157
Harte	-0.12379	-0.77349	0.525905	0.708812	0.582994	89.31295
Ye	-0.08874	-0.77257	0.595082	0.799222	0.616837	86.267
Esmail	-0.13712	-0.81169	0.537461	0.690345	0.606665	88.31534
Feng	-0.37592	-0.66872	-0.08311	0.01186	0.047308	36.49487
Zahmatkesh	0.030263	-0.56862	0.62915	0.921105	0.458542	85.20076

Table S6. Sensitivity analysis for studies assessing cholesterol levels

Table S7. Sensitivity analysis for studies assessing LDL levels

Study	Estimate	CI_lb	CI_ub	p_value	Tau2	I2
Harrison	-	-1.29994	0.138199	0.113358	0.569794	88.06166
	0.58087					
Harte	0.57903	-1.26635	0.108283	0.098702	0.540852	88.84267
Ye	0.70468	-1.35436	-0.055	0.033513	0.431989	80.88303
Esmail	0.65704	-1.34244	0.028372	0.060268	0.504801	86.14197
Feng	0.23803	-0.58207	0.106012	0.175092	0.066837	45.24189
Zahmatkesh	0.51718	-1.23872	0.204365	0.16007	0.570103	87.54376

**Table S8.** Sensitivity analysis for studies assessing FBS levels

Study	Estimate	CI_lb	CI_ub	p_value	Tau2	I2
ZELBER-SAGI	0.24824	-0.80909	0.312607	0.38566	0.38823	83.57852
Harrison	0.41123	-0.89844	0.075982	0.098066	0.272885	78.34587
Harte	- 0.16758	-0.61696	0.281803	0.46485	0.24191	78.16519
Ye	0.33762	-0.92707	0.251819	0.261591	0.42588	81.5384
Esmail	0.26809	-0.84157	0.305382	0.359529	0.407805	83.91835
Feng	0.42207	-0.89049	0.046342	0.077386	0.238131	73.96895
Zahmatkesh	- 0.17993	-0.6824	0.322533	0.482762	0.292321	78.92562